## Exam answers

HERE ARE sketches of solutions to the midterm and final exam problems.

('M1' means midterm problem 1. Get it?) M1 We have

$$\begin{split} \sum_{k} \frac{t_{k}(r)}{k+r} \, t_{n-k}(s) &= \sum_{j=0}^{m-1} \sum_{k} \frac{t_{mk+j}(r)}{mk+j+r} \, t_{n-mk-j}(s) \\ &= \sum_{j=0}^{m-1} \binom{mk+j+r}{k} \binom{n-mk-j+s}{\lfloor (n-j)/m \rfloor - k} \frac{1}{mk+j+r} \\ &= \sum_{j=0}^{m-1} \frac{1}{r+j} \binom{n+r+s}{\lfloor (n-j)/m \rfloor} = \sum_{j=0}^{m-1} \frac{t_{n-j}(r+s+j)}{r+j} \, . \end{split}$$

 $\mathbf{M2}$  (a) By induction on k; the results are easily verified when k=0 and k=1. For example,  $L_0L_{n+1}+L_{(-1)}L_n=2(F_{n+2}+F_n)-(F_{n+1}+F_{n-1})=2(F_{n+1}+2F_n)-(2F_{n+1}-F_n)=5F_n.$ 

(b) The identities of part (a) give

$$\begin{split} \text{gcd}(L_{n+k},L_n) &= \text{gcd}(F_kL_{n+1},L_n) = \text{gcd}(F_k,L_n), & \text{since } L_{n+1} \perp L_n; \\ \text{gcd}(L_{n+k},F_n) &= \text{gcd}(L_kF_{n+1},F_n) = \text{gcd}(L_k,L_n), & \text{since } F_{n+1} \perp F_n; \\ \text{gcd}(F_{n+k},L_n) &= \text{gcd}(5F_{n+k},L_n) & \text{since } 5 \perp L_n \text{ (easily verified)} \\ &= \text{gcd}(L_kL_{n+1},L_n) = \text{gcd}(L_k,L_n). \end{split}$$

Now use induction on |n/m|.

(c) We have  $gcd(2, F_n) = gcd(2, L_n) = 1 + [3 \ n]$ . There is a perfect analogy between (LL, LF, FL) and the respective functions  $F_{++}, F_{+-}, F_{-+}$  in Handout 16; so the recurrences must have an analogous outcome.

M3 There are  $1+2+\cdots+(4n+2)=\binom{4n+3}{2}$  electoral votes altogether; this is an odd number, so Big Bird needs at least  $N=\lceil \binom{4n+3}{2}/2 \rceil$ . To win in state k he needs  $k^2$  votes. If he receives more than N electoral votes he could get by with fewer. (Proof: Let k be the smallest state won. If k=1, he could lose in state 1 and still win the election. If k>1, he could more easily win in state k-1 and lose in state k.) So we can assume that he has

exactly N electoral votes, and we want to minimize  $k_1^2 + k_2^2 + \cdots + k_r^2$  such that  $k_1 + k_2 + \cdots + k_r = N$  and  $0 < k_1 < k_2 < \cdots < k_r$ .

Suppose he loses in states i and j but wins in states i-1 and j+1, where i < j. (If i=1 we can assume that he wins in state 0.) Then he could have done it easier by winning states i and j, because that needs only  $i^2+j^2 < i^2+j^2+2(j-i+1)=(i-1)^2+(j+1)^2$  votes. Therefore we can assume that there is at most one state  $\leq k_r$  that is lost. And now there's only one possibility, namely that he won all the smallest states 1, 2, ..., m except for state k, where m and k are determined by the conditions

$$1+2+\cdots+(m-1) < N \le 1+2+\cdots+m = N+k$$

For example, the smallest cases are

| Ν | states | m | k | Ν | states    | m | k |
|---|--------|---|---|---|-----------|---|---|
| 1 | 1      | 1 | 0 | 5 | 2 + 3     | 3 | 1 |
| 2 | 2      | 2 | 1 | 6 | 1 + 2 + 3 | 3 | 0 |
| 3 | 1 + 2  | 2 | 0 | 7 | 1 + 2 + 4 | 4 | 3 |
| 4 | 1 + 3  | 3 | 2 | 8 | 1 + 3 + 4 | 4 | 2 |

By exercise 3.23 we have  $m = \lfloor \sqrt{2N} + \frac{1}{2} \rfloor$ . Hence the minimum number is

$$\begin{array}{ll} p \; = \; 1^2 + 2^2 + \cdots + m^2 - k^2 \\ & = \; \frac{m(m + \frac{1}{2})(m+1)}{3} - \left(\frac{m(m+1)}{2} - N\right)^2, \end{array}$$

where m and N have already been expressed in terms of n.

In part (a) we have N=637, m=36, k=29, p=15365. The total number of penguins is 85800; hence less than 18% of the popular vote suffices.

M4 (a) The only case with a or b even is a=b=944. The other cases are  $(a,b)=(1,1987),\ (3,1985),\ \dots,\ (1989,-1)$ . (Notice the tricky ending.)

- (b) If  $f_m = f_n$  then  $m + 1998 = f_{f_m} = f_{f_n} = n + 1988$ .
- (c) We have  $f_{n+1988} = f_{f_{n}} = f_{n} + 1988$ .
- (d)  $a_0 = f_0$ ,  $a_1 = f_1 1$ , ...,  $a_{1987} = f_{1987} 1987$ ; use induction on n.
- (e) If  $n \in A$  then  $f_n < 1988$  and  $f_{f_n} = n + 1988 \geqslant 1988$ , hence  $f_n \in B$ . Hence  $\#(A) \leqslant \#(B)$ .
- (f) If  $n \in B$  then  $1988 \leqslant f_n = n' + 1988$  for some n'. We have  $n + 1988 = f_{f_n} = f_{n'+1988} = f_{n'} + 1988$ ; hence  $f_{n'} = n < 1988$ . Now if n' were  $\geqslant 1988$ , we'd have  $f_{n'} \geqslant 1988$  by part (c). Hence  $n' \in A$ ; we have proved that  $n \in B \Longrightarrow f_n 1988 \in A$ . Hence  $\#(B) \leqslant \#(A)$ .

(g) Let the elements of A be  $a_1, a_2, \ldots, a_{994}$  and let the corresponding elements of B be  $b_1 = f_{a_1}, \ldots, b_{994} = f_{a_{994}}$ . Then

$$f_n \ = \ \begin{cases} n + b_k - a_k, & \text{if } n \text{ mod } 1988 = a_k; \\ n + a_k + 1988 - b_k, & \text{if } n \text{ mod } 1988 = b_k. \end{cases}$$

So 
$$\sum_{k=1}^{1988} f_k = (0+1+\dots+1987) + 994 \cdot 1988.$$

So  $\sum_{k=1}^{1988} f_k = (0+1+\cdots+1987)+994\cdot 1988$ . (h) The number of ways to choose A is  $\binom{1988}{994}$ , and the number of ways to assign b's to a's is 994!; hence the number of possibilities is 1988!/994!.

 $\begin{array}{ll} \mathbf{F1} & A_{\mathfrak{m},\mathfrak{n}}(z) = \big(z\big(\frac{1+z}{2}\big)\dots\big(\frac{\mathfrak{n}-1+z}{\mathfrak{n}}\big)\big)^{\mathfrak{m}} = \big(\frac{z+\mathfrak{n}-1}{\mathfrak{n}}\big)^{\mathfrak{m}}. & \text{Hence Mean } A_{\mathfrak{m},\mathfrak{n}} = \mathfrak{m}\cdot \text{Mean } A_{1,\mathfrak{n}} = \mathfrak{m} B_{\mathfrak{n}}, & \text{Hence Mean } A_{\mathfrak{m},\mathfrak{n}} = \mathfrak{m}\cdot \text{Var } A_{1,\mathfrak{n}} = \mathfrak{m}\sum_{k=1}^{\mathfrak{n}} \frac{1}{\mathfrak{n}}\big(\frac{\mathfrak{n}-1}{\mathfrak{n}}\big) = \mathbb{E}_{\mathfrak{m}} B_{\mathfrak{m}} B_{\mathfrak$  $\mathfrak{m}(H_{\mathfrak{n}}-H_{\mathfrak{n}}^{(2)})$ . The variance can also be expressed as  $\mathfrak{m}(\frac{2}{\mathfrak{n}!}{n! \choose 3}+H_{\mathfrak{n}}-H_{\mathfrak{n}}^2)$ . Furthermore.

$$B_{m,n}(z) = z \left(\frac{2^m - 1 + z}{2^m}\right) \left(\frac{3^m - 1 + z}{3^m}\right) \dots \left(\frac{n^m - 1 + z}{n^m}\right);$$

hence Mean  $B_{\mathfrak{m},\mathfrak{n}}=H_{\mathfrak{n}}^{(\mathfrak{m})},\, Var(B_{\mathfrak{m},\mathfrak{n}})=H_{\mathfrak{n}}^{(\mathfrak{m})}-H_{\mathfrak{n}}^{(2\mathfrak{m})}.$  And

$$\begin{array}{ll} C_{\mathfrak{m},\mathfrak{n}}(z) \; = \; z \left( \frac{1 + (2^{\mathfrak{m}} - 1)z}{2^{\mathfrak{m}}} \right) \left( \frac{2^{\mathfrak{m}} + (3^{\mathfrak{m}} - 2^{\mathfrak{m}})z}{3^{\mathfrak{m}}} \right) \\ & \qquad \cdots \left( \frac{(n-1)^{\mathfrak{m}} + (n^{\mathfrak{m}} - (n-1)^{\mathfrak{m}})z}{n^{\mathfrak{m}}} \right) \; ; \end{array}$$

hence Mean  $C_{\mathfrak{m},\mathfrak{n}}=\mathfrak{n}-\left(\left(\frac{0}{1}\right)^{\mathfrak{m}}+\left(\frac{1}{2}\right)^{\mathfrak{m}}+\cdots\left(\frac{\mathfrak{n}-1}{\mathfrak{n}}\right)^{\mathfrak{m}}\right)=\mathfrak{n}-S_{\mathfrak{m},\mathfrak{n}}.$  Finally,  $Var C_{m,n} = S_{m,n} - S_{2m,n}.$ 

**F2** (a) Each term is at most  $(1-n^{-4/5})^{cn} < \exp(-n^{-4/5})^{cn} = \exp(-cn^{1/5})$ , which is "superpolynomially small," i.e.,  $O(n^{-m})$  for all m. So the whole sum is superpolynomially small.

- (b)  $\ln(1-1/k) = -1/k 1/2k^2 O(k^{-3})$ , and  $\operatorname{cnO}(k^{-3}) = O(n^{-7/5})$  in this range.
- (c)  $\exp(-\alpha + O(n^{-7/5})) = e^{-\alpha} + O(n^{-7/5})$  when  $\alpha \ge 0$ , and there are at most n terms in the sum.
- (d) Let  $f(x) = \exp(-cn/x cn/2x^2)$ . Then when  $\lfloor n^{4/5} \rfloor \leqslant x \leqslant n$  we have

$$f'(x) = \left(\frac{cn}{x^2} + \frac{cn}{x^3}\right) f(x), \qquad f''(x) = O(n^{-6/5}).$$

Applying (9.67) and (9.68) with m = 2 gives a remainder  $R_2 = O(n^{-1/5})$ , because the integrand is  $O(n^{-6/5})$  and the range of integration is O(n). Furthermore f(x) and f'(x) are superpolynomially small when  $x = \lfloor n^{4/5} \rfloor$ ; and

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 $f'(n) = O(n^{-1})$ . Therefore

$$S(cn,n) = \int_{n^{4/5}}^{n} f(x) dx + \frac{1}{2} f(n) + O(n^{-1/5}),$$

and  $f(n) = e^{-c} + O(n^{-1})$ . The substitution x = n/y converts the remaining integral into

(We did not expect everybody to solve all the problems on the final.)

$$n \int_{1}^{n^{1/5}} e^{-cy} \left( 1 - \frac{cy^{2}}{2n} + O(n^{-7/5}) \right) \frac{dy}{y^{2}}$$

$$= n \int_{1}^{\infty} e^{-cy} \frac{dy}{y^{2}} - \frac{c}{2} \int_{1}^{\infty} e^{-cy} dy + O(n^{-2/5})$$

$$= cn \int_{c}^{\infty} e^{-t} \frac{dt}{t^{2}} - \frac{e^{-c}}{2} + O(n^{-2/5}).$$

(Here we can integrate all the way to  $\infty$ , because

$$\int_{n^{1/5}}^{\infty} e^{-ct} \frac{dt}{t^2} \leqslant \int_{n^{1/5}}^{\infty} e^{-ct} \frac{dt}{n^{2/5}}$$

is superpolynomially small.) The integral that still remains is  $e^{-c}-\mathsf{E}_1(c)$ . Hence the answer comes to

$$a(c) = e^{-c} - cE_1(c), \quad b(c) = 0.$$

The same method gives  $O(n^{-1+4\epsilon})$  if we replace 4/5 by  $1-\epsilon$ .

**F3** (a) 
$$T_r(z) = \sum_n {n+r \choose \lfloor n/2 \rfloor} z^n = \sum_n {2n+r \choose n} z^{2n} + \sum_n {2n+1+r \choose n} z^{2n+1}$$
, which is

$$\frac{\mathcal{B}_2(z^2)^{\mathrm{r}}}{\sqrt{1-4z^2}}\left(1+\frac{1-\sqrt{1-4z^2}}{2z}\right),\qquad \mathcal{B}_2(z^2) = 1+\frac{1-\sqrt{1-4z^2}}{2z},$$

by (5.78) and (5.82).

- (b) The sum is  $[z^n] T_r(z) T_{s-r}(z) = [z^n] T_0(z) T_s(z)$ .
- (c) The generating function for the sum is

$$T_0(z)^2 = \left(\frac{1}{\sqrt{1-4z^2}} \left(1 + \frac{1-\sqrt{1-4z^2}}{2z}\right)\right)^2$$

$$= \frac{1}{4z^2(1-4z^2)} \left(4z^2 + 4z\left(1 - \sqrt{1-4z^2} + 1 - 2\sqrt{1-4z^2} + 1 - 4z^2\right)\right)$$

$$= \frac{1}{2z^2(1-4z^2)} \left(2z + 1 - (2z+1)\sqrt{1-4z^2}\right)$$

$$\begin{split} &= \frac{1}{2z^{2}(1-2z)} - \frac{2z+1}{2z^{2}\sqrt{1-4z^{2}}} \\ &= \sum_{n \geq 0} \frac{2^{n}z^{n}}{2z^{2}} - \sum_{n \geq 0} {2n \choose n} z^{2n-1} - \frac{1}{2} \sum_{n \geq 0} {2n \choose n} z^{2n-2} \\ &= \sum_{n \geq 0} 2^{n+1}z^{n} - \sum_{n \geq 0} {2n+2 \choose n+1} z^{2n+1} - \frac{1}{2} \sum_{n \geq 0} {2n+1 \choose n} z^{2n}; \end{split}$$

so the sum is  $2^{n+1} - t_{n+1}(0)$ .

- F4 (a) The total number Y of cliques is the sum of  $\binom{n}{r}$  random variables  $Y_k$ , where each  $Y_k$  is 1 if a particular set of r students forms a clique, and  $Y_k = 0$  otherwise. The probability that a particular set of r students forms a clique is  $p^{\binom{r}{2}}$ ; hence the expected number of cliques is  $\binom{n}{r}p^{\binom{r}{2}}$ . Similarly, the expected number of claques is  $\binom{n}{s}(1-p)^{\binom{s}{2}}$ .
- (b) Consider the random variable  $X = [there is at least one r-clique or s-claque]; and let Y and Z denote the number of r-cliques and s-claques. Then <math>X \leq Y + Z$ . So EX = 1 implies that  $E(Y + Z) \geqslant 1$ .
- (c) Let  $p=1/\varphi^2$ . (It turns out that no other value of p will work in the following argument.) We wish to show that

$$\binom{n}{r}p^{\binom{r}{2}}+\binom{n}{2r}(1-p)^{\binom{2r}{2}}\ <\ 1\,,\qquad \text{where } n=rF_{r-2},$$

for all sufficiently large r. But the left-hand side is in fact exponentially small as  $r \to \infty$ :

$$\begin{split} \ln \left( \binom{n}{r} p^{\binom{r}{2}} \right) \; &< \; \ln \left( \frac{n^r}{r!} \, \varphi^{r(r-1)} \right) \\ &= \; r \left( \ln n - (r-1) \ln \varphi - \ln r + 1 \right) + O(\log r) \\ &= \; r \left( \ln F_{r-2} - (r-1) \ln \varphi + 1 \right) + O(\log r) \\ &= \; r \left( -\ln \varphi - \ln \sqrt{5} + 1 \right) + O(\log r); \\ \ln \left( \binom{n}{2r} (1-p)^{\binom{2r}{2}} \right) \; &< \; \ln \frac{n^{2r}}{(2r)!} \varphi^{r(2r-1)} \\ &= \; r \left( 2 \ln n - (2r-1) \ln \varphi - 2 \ln 2r + 2 \right) + O(\log r) \\ &= \; r \left( 2 \ln F_{r-2} - (r-1) \ln \varphi - 2 \ln 2 + 1 \right) + O(\log r) \\ &= \; r \left( -\ln \varphi - \ln \sqrt{5} - 2 \ln 2 + 2 \right) + O(\log r). \end{split}$$

(The stated inequality turns out to be true for all  $r \ge 2$ .)

**F5** (a)  $b_{n+1} = b_n + \frac{1}{n+1} \sum_{k=1}^{n-1} b_k b_{n-k} \binom{n}{k}^{-1}$ , hence  $b_1 \leqslant b_2 \leqslant \cdots$  and we can get an upper bound  $b_n + \frac{1}{n+1} \sum_{k=1}^{n-1} b_n^2 \binom{n}{k}^{-1}$ . The stated inequality is

(This fact should have been stated as a hint.)

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obvious when n = 1 or n = 2, and for  $n \ge 3$  we have

$$\sum_{k=1}^{n-1} \binom{n}{k}^{-1} \; \leqslant \; \frac{1}{n} + \sum_{k=2}^{n-2} \binom{n}{k}^{-1} + \frac{1}{n} \; = \; \frac{2}{n} + \frac{2(n-3)}{n(n-1)} \; < \; \frac{4}{n} \, .$$

(b) If  $b_n\leqslant c\sqrt{n}$  then  $b_{n+1}\leqslant c\sqrt{n}+4c^2/(n+1),$  and this is at most  $c\sqrt{n+1}$  if  $4c\leqslant (n+1)\big(\sqrt{n+1}-\sqrt{n}\big).$  The condition holds when  $n\geqslant 16$  and c=0.53; it also holds for c=1 when  $n\geqslant 63.$  Thus,  $b_n=O\big(\sqrt{n}\big).$  Hence  $b_{n+1}=b_n+O\big((\log n)^2/n^2\big),$  and we have  $b_n=O(\log n).$  Hence  $b_{n+1}=b_n+O\big((\log n)^2/n^2\big),$  and we have  $b_n=O(1).$  Finally, since the sequence  $\langle b_n\rangle$  is increasing and bounded, it must approach a limit  $\alpha.$ 

(A computer was helpful here.)

(c) Since  $b_n \to \alpha$ , we have

$$\alpha - b_n = \sum_{k \geqslant n} (b_{k+1} - b_k) \leqslant \sum_{k \geqslant n} \frac{4b_k^2}{k(k+1)} = O(1/n).$$

(d) Here's how to solve the recurrence with generating functions. Let  $A(z) = a_1 z + a_2 z^2 + \cdots = \vartheta U(z)$ , where  $\vartheta = z \frac{d}{dz}$ , and let  $C(z) = e^{U(z)}$ . Then

(This solution, found independently by Alon L., San-joy M., Steven P., and Heping Z., avoids the need for the lemma presented in class.)

$$A = z\vartheta A + zA^2 + zA + z,$$
  
 $\vartheta C = AC, \qquad \vartheta^2 C = (\vartheta A)C + A^2C;$ 

so

$$\vartheta C = z(\vartheta - \omega)(\vartheta - \omega^2)C$$

where  $\omega = (-1 + \sqrt{3}i)/2$ . Hence  $C(z) = F(-\omega, -\omega^2; z)$  by (5.119); and

 $[z^{\mathbf{n}}] C(z) = \frac{(-\omega)^{\overline{\mathbf{n}}} (-\omega^2)^{\overline{\mathbf{n}}}}{\mathbf{n}!} = \frac{\Gamma(\mathbf{n} - \omega)\Gamma(\mathbf{n} - \omega^2)}{\Gamma(-\omega)\Gamma(-\omega^2) \mathbf{n}!} \sim \frac{\Gamma(\mathbf{n})}{\Gamma(-\omega)\Gamma(-\omega^2)}$ 

This bonus problem was not as easy as the one on the midterm.

because

$$\frac{\Gamma(n+a)\Gamma(n+b)}{\Gamma(n+c)\Gamma(n+d)} \sim 1 \quad \text{when } a+b=c+d.$$

These generating functions are divergent; indeed, they diverge so fast, we have  $c_n \sim u_n = \frac{1}{n} a_n$ . Hence

$$\alpha \; = \; \frac{1}{\Gamma(-\omega)\Gamma(-\omega^2)} \; = \; \frac{\cosh\left(\pi\sqrt{3}/2\right)}{\pi} \, , \label{eq:alpha}$$

by (5.96) and (5.97).

Merry Christmas and Happy New Year to all!